

ECE316- Probability and Random Variables Winter 2007
In-term Exam 1

Time: 1 Hour 15 minutes

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INSTRUCTIONS

- Enter your name, student ID number, e-mail address and sign in the space provided at the bottom of this page.
- The exam is 1 hour and 15 mins long.
- DO ALL QUESTIONS. The weightage of each question is indicated.
- Some useful results and formulae are given on the last page.
- You are **not** allowed the use of calculators or crib sheets.
- Unless specifically mentioned to the contrary show all relevant work in the space provided. You may use the extra blank pages and the back of each page if necessary.

Name:

Student ID #:

E-mail :

Signature:

SOLUTION

PART I

(20 %)

These are short problems. For problems involving TRUE or FALSE answers it is not necessary to justify your answers. Only correct answers count- no partial credit.

1. (4pts) Let A and B be two events defined on (Ω, \mathcal{F}, P) : Then:

a) $\Pr(A) + \Pr(B) \leq 1$. *Consider 2 events with $P(A) > 0.5$ and $P(B) > 0.5$*

Answer: True False *Then $P(A) + P(B) > 1$*

b) $\Pr(A \cup B) \geq \max\{\Pr(A), \Pr(B)\}$

Answer: True False
Since both $A \subset A \cup B$ and $B \subset A \cup B$

2. (6pts) Suppose A, B, C are three events defined on the same probability space.

a) The events A and B are independent if $A \cap (B^c) = \emptyset$

Answer: True False *If $A \cap B^c = \emptyset$ then A and B^c are dependent and so are A and B*

b) If A and B are independent, and B and C are independent. Then A and C are independent.

Answer: True False *This info tells you nothing about A and C*

c) Suppose (A, B) , (B, C) and (C, A) are pairs of independent events, then A, B, C are independent.

Answer: True False *Pairwise independence is not enough. See class notes. - example*

3. (4 pts) Let A and B be two events with $\Pr(A) > 0$ and $\Pr(B) > 0$.

Show that $\Pr(A|B) \geq \Pr(A)$ then $\Pr(B|A) \geq \Pr(B)$.

$$P(A \cap B) = P(A|B) \cdot P(B) = P(B|A) P(A)$$

$$\text{So, if } P(A|B) \geq P(A) \Rightarrow P(A \cap B) \geq P(A)P(B)$$

$$\text{Hence } P(B|A) = \frac{P(A \cap B)}{P(A)} \geq \frac{P(A)P(B)}{P(A)} = P(B)$$

4. (6pts) An event B is said to suggest event A if $\Pr(A|B) \geq \Pr(A)$ and does not suggest it if $\Pr(A|B) < \Pr(A)$. Show that if $0 < \Pr(B) < 1$ then the following is true:

B suggests A if and only if B^c does not suggest A .

Law of total prob.

$$P(A) = P(A|B)P(B) + P(A|B^c)P(B^c) \quad (i)$$

On the other hand

$$P(A) = P(A) [P(B) + P(B^c)] \quad (ii)$$

So, $(i) - (ii) = 0$

$$P(B) [P(A|B) - P(A)] = P(B^c) [P(A) - P(A|B^c)]$$

So clearly if $P(A|B) \geq P(A)$ (B suggests A)

$\Leftrightarrow P(A) \leq P(A|B^c)$ or B^c does not suggest A .

PART II

For the problems in this section show all calculations in the space provided or at the back of the page. Justify all your answers.

Problem 1: (16%)

Two numbers X and Y are selected at random between -1 and 1 . Let the events A , B , C and D be defined as follows: $A = \{X > 0\}$, $B = \{Y > 0\}$, $C = \{X < -1/2 \text{ and } Y > 1/2\}$ and $D = \{X < 0 \text{ and } Y < 0\} \cup \{X > 0 \text{ and } Y > 0\}$.

a) Are A and B independent? Why?

X and Y are independent (they are selected at random)
so A and B are independent

b) Are A and C independent? Why?

A and C : No.

$$A \cap C = \{x > 0\} \cap \left\{ \left\{ x < -\frac{1}{2} \right\} \cap \left\{ y > \frac{1}{2} \right\} \right\}$$

$$= \emptyset$$

so A and C cannot be independent

c) Are A, B and C independent? Why?

$(A \cap B \cap C)$? No. Since $A \cap B \cap C = \emptyset$
as in part b)

d) Are A, B and D independent? Why? Find $\Pr(A \cap B \cap D)$.

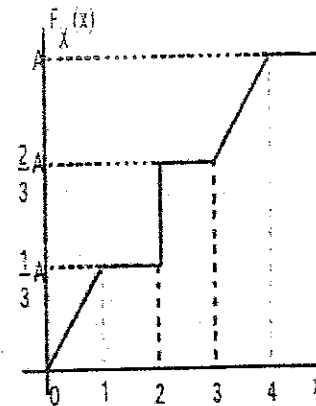
$$\begin{aligned} A \cap B \cap D &= \{x > 0\} \cap \{y > 0\} \cap \left\{ \{x < 0\} \cap \{y < 0\} \right. \\ &\quad \left. \cup \{x > 0\} \cap \{y > 0\} \right\} \\ &= \{x > 0\} \cap \{y > 0\} = A \cap B \end{aligned}$$

So A, B and D would be independent iff $\Pr(D) = 1$,
which is clearly not the case. So A, B and D are
not independent

$$\begin{aligned} \Pr(A \cap B \cap D) &= \Pr(\{x > 0\} \cap \{y > 0\}) = \frac{1}{2} \times \frac{1}{2} = \frac{1}{4} \\ &\text{(since } X \text{ and } Y \text{ are uniform on } [-1, 1]) \end{aligned}$$

Problem 3: (24%)

A random variable X has the following cumulative probability distribution function (c.d.f) $F(x)$ shown:

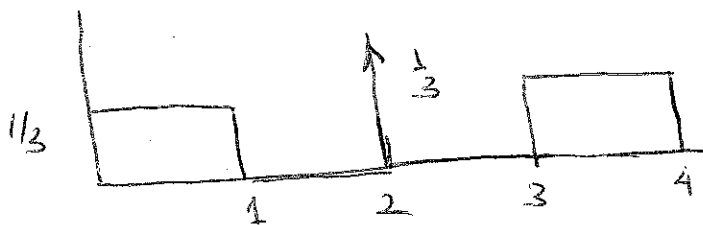


1. What is the value of A ? Why?

$$A = 1 \quad \text{since} \quad F(x) \rightarrow 1 \quad \text{as} \quad x \rightarrow \infty$$

2. Sketch the probability density function.

$$P_X(x) = F'(x)$$



3. Find $E[X]$.

$$\begin{aligned} E[X] &= \frac{1}{3} \int_0^1 x \, dx + \frac{2}{3} + \frac{1}{3} \int_3^4 x \, dx \\ &= \frac{1}{6} + \frac{2}{3} + \frac{1}{3} \left[\frac{16}{2} - \frac{9}{2} \right] \\ &= \frac{1}{6} + \frac{4}{6} + \frac{7}{6} = \frac{12}{6} = 2 \end{aligned}$$

4. What is the probability that $X < 2$?

$$P\{X < 2\} = \frac{1}{3} \quad (\text{Right continuous so value to the left of } 2)$$

5. Find the variance of X .

$$\text{Variance of } X = E[X^2] - E[X]^2$$

$$E[X^2] = \frac{1}{3} \cdot \frac{1}{3} + \frac{4}{3} + \frac{1}{3} \int_3^9 x^2 dx$$

$$= \frac{1}{9} + \frac{4}{3} + \frac{64 - 27}{9}$$

$$= \frac{1 + 12 + 37}{9} = \frac{50}{9}$$

$$\text{Var } X = \frac{50}{9} - 4 = \frac{14}{9}$$

Problem 4: (16%)

Let X be a r.v. with probability mass function (p.m.f) given by:

$$p_k = \Pr(X = k) = \frac{k^2}{C}, \quad k = 0, \pm 1, \pm 2, \pm 3 \\ = 0, \text{ otherwise}$$

1. Find C and $E[X]$.

$$\sum p_k = 1 = \frac{2}{C} [1 + 4 + 9] \Rightarrow C = 28$$

$$E[X] = 0 \text{ by symmetry}$$

2. Find the p.m.f. of the random variable Z defined as:

$$Z = (X - E[X])^2$$

Now $Z = X^2$ so the p.m.f of Z
 Z takes values 1, 4, 9

$$P\{Z=0\} = 0$$

$$P\{Z=1\} = \frac{1}{28} + \frac{1}{28} = \frac{1}{14}$$

$$P\{Z=4\} = 2 \times \frac{4}{28} = \frac{8}{28} = \frac{2}{7}$$

$$P\{Z=9\} = 2 \times \frac{9}{28} = \frac{18}{28} = \frac{9}{14}$$

3. Find the variance of X .

By defn. of Z

$$\text{Var } X = E[Z^2]$$

$$= \frac{1}{14} + \frac{8}{7} + \frac{81}{14}$$

$$= \frac{1 + 16 + 81}{14} = \frac{98}{14} = 7$$

Problem 5: (24%)

Let X and Y be independent geometric random variables with parameter p i.e.:

$$\Pr(X = n) = \Pr(Y = n) = (1-p)^{n-1}p, \quad n \geq 1$$

1. Find the moment generating function of X and hence $E[X]$.

$$\begin{aligned}\Phi(z) &= \sum_{k=1}^{\infty} z^k (1-p)^{k-1} p \\ &= pz \sum_{k=1}^{\infty} [z(1-p)]^{k-1} \\ &= pz \sum_{k=0}^{\infty} [z(1-p)]^k \\ &= \frac{pz}{1-z(1-p)}.\end{aligned}$$

$$\begin{aligned}E[X] &= \left. \frac{d}{dz} \Phi(z) \right|_{z=1} = \left. \frac{p}{1-z(1-p)} \right|_{z=1} \\ &\quad + \left. \frac{pz(1-p)}{(1-z(1-p))^2} \right|_{z=1} \\ &= \frac{p}{p} + \frac{p(1-p)}{p^2} \\ &= 1 + \frac{(1-p)p}{p^2} \\ &= \frac{p}{p^2} = \frac{1}{p}.\end{aligned}$$

(see class notes)
HW-3 also.

2. Show that

$$\Pr(X > n+m | X > m) = \Pr(X > n)$$

$$P\{X > n+m | X > m\} = \frac{P\{(X > n+m) \cap (X > m)\}}{P\{X > m\}} = \frac{P\{X > n+m\}}{P\{X > m\}}$$

$$P\{X > k\} = P \sum_{k+1}^{\infty} (1-p)^{n-1} = P \sum_{k+1}^{\infty} (1-p)^n = (1-p)^k$$

$$\text{So } P\{X > n+m | X > m\} = \frac{(1-p)^{n+m}}{(1-p)^m} = (1-p)^n = P\{X > n\}$$

(See HW-3)

3. Find $\Pr(X+Y=n)$

Now

$$P\{X+Y=n\} = \sum_{k=1}^{n-1} P\{X=k, Y=n-k\} = \sum_{k=1}^{n-1} P_k P_{n-k}$$

$$= \sum_{k=1}^{n-1} (1-p)^{k-1} p (1-p)^{n-k-1} p \quad (\text{see lecture notes})$$

$$= p^2 (1-p)^{n-2} \sum_{k=1}^{n-1} 1$$

$$= (n-1) p^2 (1-p)^{n-2}$$

4. Show that

$$\Pr(X = k | X + Y = n + 1) = \frac{1}{n}, \quad k = 1, 2, \dots, n$$

$$P\{X = k | X + Y = n + 1\} = \frac{P\{X = k, X + Y = n + 1\}}{P\{X + Y = n + 1\}}$$

$$= \frac{P\{X = k, Y = n + 1 - k\}}{P\{X + Y = n + 1\}}$$

$$= \frac{P (1-P)^{k-1} P (1-P)^{n-k}}{n P^2 (1-P)^{n-1}}$$

$$= \frac{1}{n}$$

Note this does not depend on k

$$\text{so } P\{X = k | X + Y = n + 1\} = \frac{1}{n}, \quad k = 1, 2, \dots, n.$$

Some useful formulae and notes

Abbreviations used: Pr. = Probability, r.v. = Random Variable, c.d.f. = Cumulative Distribution Function (denoted $F(x)$), p.d.f = Probability density function (for continuous distributions) (usually denoted $p(x)$), p.m.f= Probability mass function (for discrete distributions) (usually denoted $p_k = \Pr(X = x_k)$).

For right continuous functions $G(x)$ the derivative is defined from the right i.e.:

$$g(x) = \frac{dG(x)}{dx} = \lim_{\Delta \rightarrow 0} \frac{G(x + \Delta) - G(x)}{\Delta}$$

when it exists.

Some elementary formulae

: If X is a discrete-r.v. then the k - th moment of X :

$$\mathbf{E}[X^k] = \sum_n x_n^k p_n$$

where $p_n = \Pr(X = x_n)$.

If X is continuous then:

$$\mathbf{E}[X^k] = \int_{\mathfrak{R}} x^k dF(x) = \int_{\mathfrak{R}} x^k p(x) dx$$

where $p(x) = \frac{dF(x)}{dx}$ is the p.d.f. and $F(x)$ is the c.d.f.

Some other elementary formulae

Given a collection of sets $\{A_i\}$ then De-Morgan's law states:

$$\left(\bigcup_{k=1}^n A_k \right)^c = \bigcap_{k=1}^n A_k^c$$

where $A_k^c = \Omega - A_k$ (the complement of A_k).

$$e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$$

$$\sum_{k=0}^N r^k = \frac{1 - r^{N+1}}{1 - r}$$

$$\sum_{k=0}^{\infty} r^k = \frac{1}{1 - r} \quad \text{if } |r| < 1$$