

1. (a)  $\Pr\{X(\omega) \leq 1\} = F(1) = 1 - e^{-2}$   
 $\Pr\{X(\omega) > 2\} = 1 - F(2) = 1 - (1 - e^{-4}) = e^{-4}$   
 $\Pr\{X(\omega) = 2\} = 0$  since  $X(\omega)$  is a continuous R.V.

(b) pdf.  $p_X(x)$  of  $X(\omega)$   

$$p_X(x) = \frac{d}{dx} \{F_X(x)\} = \begin{cases} 2e^{-2x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

(c)  $\Pr\{Y(\omega) = 0\} = \Pr\{X(\omega) \leq 2\} = 1 - e^{-4}$   
 $\Pr\{Y(\omega) = 1\} = \Pr\{X(\omega) > 2\} = e^{-4}$   
 $P_Y(y) = (1 - e^{-4})\delta(y) + e^{-4}\delta(y-1)$

2. (a)  $X(\omega) \geq 0$   
 By the def. of  $E(X(\omega))$

$$E(X(\omega)) = \int_0^{\infty} x dF(x)$$

Since  $dF_X(x) = -d(1 - F_X(x))$

$$E(X(\omega)) = -\int_0^{\infty} x d(1 - F_X(x))$$

Use Integration by parts.

$$= -x(1 - F(x)) \Big|_0^{\infty} + \int_0^{\infty} (1 - F(x)) dx$$

Since  $E|X| < \infty$

$$\lim_{x \rightarrow \infty} x(1 - F(x)) = 0$$

$$= 0 + \int_0^{\infty} (1 - F(x)) dx = \int_0^{\infty} (1 - F(x)) dx$$

0.

2. (b)

$$E(X(\omega)) = \int_{-\infty}^{\infty} x dF(x)$$

$$= \int_{-\infty}^0 x dF(x) + \int_0^{\infty} x dF(x)$$

use integration by parts then,

$$= x F(x) \Big|_{-\infty}^0 - \int_{-\infty}^0 F(x) dx + \int_0^{\infty} (1-F(x)) dx$$

Since if  $E|X| < \infty$

$$\lim_{x \rightarrow \infty} x F(x) = 0$$

$$= \int_0^{\infty} (1-F(x)) dx - \int_{-\infty}^0 F(x) dx$$

□

3. Assume  $A = \{X+Y > z\}$   $B = \{X > z\}$

$$\text{Use } P(A) = P(A|B)P(B) + P(A|B^c)P(B^c)$$

$$\therefore P(X+Y > z) = \underbrace{P(X+Y > z | X > z)}_{= 1 \text{ since } Y \text{ is non-negative.}} P(X > z) + P(X+Y > z | X \leq z) P(X \leq z)$$

$$= P(X > z) + P(X+Y > z \geq X)$$

$$\int_0^{\infty} P(X+Y > z \geq X) dz = \int_0^{\infty} (P(X+Y > z) - P(X > z)) dz$$

from the result of 2. (a)

$$= E(X+Y) - E(X) = E(Y(\omega))$$

□

4. (a)  $X(\omega)$  takes only two values  $(-1, 1)$

$$P_z(z) = P(z|X=-1) \cdot P(X=-1) + P(z|X=1) \cdot P(X=1)$$

$$\begin{cases} P_{z|X}(z|X=-1) = P_{Y|X}(z+1|X=-1) = P_Y(z+1) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(z+1)^2}{2\sigma^2}} \\ P_{z|X}(z|X=1) = P_{Y|X}(z-1|X=1) = P_Y(z-1) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(z-1)^2}{2\sigma^2}} \end{cases}$$

$$= \frac{1}{2} \frac{1}{\sqrt{2\pi}\sigma^2} \left\{ e^{-\frac{(z+1)^2}{2\sigma^2}} + e^{-\frac{(z-1)^2}{2\sigma^2}} \right\}$$

(b) Part a.

(c)  $E(W) = E(XY) = E(X)E(Y) = 0$   
ind.

$$\text{Var}(W) = E(X^2Y^2) = E(X^2)E(Y^2) = E(Y^2) = \sigma^2$$

Uncorrelated?

$$E(WY) = E(X \cdot Y^3) = E(X) \cdot E(Y^3) = 0$$

$$E(W) \cdot E(Y) = 0$$

$\therefore E(WY) = E(W)E(Y) \Rightarrow$  uncorrelated.

Independence?

$$E(W^2Y^4) = E(X^2Y^4) = E(X^2)E(Y^4) = 1 \cdot (3 \cdot \sigma^4) = 3\sigma^4$$

$$E(W^2)E(Y^4) = \sigma^2 \cdot \sigma^4 = \sigma^6$$

$\therefore E(W^2Y^4) \neq E(W^2)E(Y^4) \therefore W \& Y$  are not independent

5.  $X_1(\omega), X_2(\omega)$

$$P(X_1, X_2) = \frac{1}{2\pi} e^{-\frac{X_1^2 + X_2^2}{2}}$$

$\rightarrow$  jointly Gaussian with mean 0, variance 1.

$$Y(\omega) = \sqrt{X_1^2(\omega) + X_2^2(\omega)}$$

Change the coordinate.

$$\cos \phi = \frac{X_1}{\sqrt{X_1^2 + X_2^2}} \quad \sin \phi = \frac{X_2}{\sqrt{X_1^2 + X_2^2}}$$

Let  $X_1 = Y \cos \phi, X_2 = Y \sin \phi$

$$P(Y, \phi) = \begin{vmatrix} \cos \phi & -Y \sin \phi \\ \sin \phi & Y \cos \phi \end{vmatrix} \frac{1}{2\pi} e^{-\frac{Y^2}{2}} = \frac{1}{2\pi} Y e^{-\frac{Y^2}{2}} \quad (\text{for } Y \geq 0)$$

$$f(x, \phi) = \frac{1}{2\pi} \cdot y e^{-\frac{y^2}{2}} = f_Y(y) \cdot f_\phi(\phi)$$

$$\therefore P_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} : y \geq 0$$

$$P_\phi(\phi) = \frac{1}{2\pi} : |\phi| < \pi$$

$\therefore Y$  &  $\phi$  are independent.

$$E[X_1 | Y] = E(Y \cos \phi | Y) = Y E(\cos \phi) = 0 \quad \text{since } E(\cos \phi) = 0 \quad \square$$

$$\begin{aligned} 6. \quad E(X_1 | X_2) &= E(Y - \alpha X_2 | X_2) = E(Y | X_2) - E(\alpha X_2 | X_2) \\ &= E(Y) - \alpha X_2 = E(X_1 + \alpha X_2) - \alpha X_2 \\ &= \cancel{E(X_1)} + \alpha \cancel{E(X_2)} - \alpha X_2 = -\alpha X_2 \quad \square \end{aligned}$$

7. Without loss of generality let us take  $E[X] = 0$

$$E - X \leq (E - X) \mathbb{1}_{(E > X)}$$

Hence taking Expectations

$$E \leq E[(E - X) \mathbb{1}_{(E > X)}]$$

$$\text{So } E^2 \leq [E[(E - X) \mathbb{1}_{(E > X)}]]^2$$

$$< E[(E - X)^2 \mathbb{1}_{\{X \leq E\}}]$$

$$\Rightarrow P\{X \leq E\} \geq \frac{E^2}{E^2 + \text{Var } X}$$

$$\text{So } P\{X > E\} = 1 - P\{X \leq E\} \leq 1 - \frac{E^2}{E^2 + \text{Var } X} = \frac{\text{Var } X}{E^2 + \text{Var } X}$$